## Chapter 10: Inferences based on two samples

## **MATH 450**

### November 14<sup>th</sup>, 2017

MATH 450 Chapter 10: Inferences based on two samples

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10.1 Difference between two population means

- z-test
- confidence intervals
- 10.2 The two-sample t test and confidence interval
- 10.3 Analysis of paired data

### Example

Let  $\mu_1$  and  $\mu_2$  denote true average decrease in cholesterol for two drugs. From two independent samples  $X_1, X_2, \ldots, X_m$  and  $Y_1, Y_2, \ldots, Y_n$ , we want to test:

 $H_0: \mu_1 = \mu_2$  $H_a: \mu_1 \neq \mu_2$ 

• This week: independent samples

## Assumption

- X<sub>1</sub>, X<sub>2</sub>,..., X<sub>m</sub> is a random sample from a population with mean μ<sub>1</sub> and variance σ<sub>1</sub><sup>2</sup>.
- **2**  $Y_1, Y_2, \ldots, Y_n$  is a random sample from a population with mean  $\mu_2$  and variance  $\sigma_2^2$ .
- **(3)** The X and Y samples are independent of each other.
  - Next week: paired-sample test

#### Problem

Assume that

- X<sub>1</sub>, X<sub>2</sub>,..., X<sub>m</sub> is a random sample from a population with mean μ<sub>1</sub> and variance σ<sub>1</sub><sup>2</sup>.
- Y<sub>1</sub>, Y<sub>2</sub>,..., Y<sub>n</sub> is a random sample from a population with mean μ<sub>2</sub> and variance σ<sub>2</sub><sup>2</sup>.
- The X and Y samples are independent of each other.

Compute (in terms of  $\mu_1, \mu_2, \sigma_1, \sigma_2, m, n$ )

(a) 
$$E[\bar{X} - \bar{Y}]$$
  
(b)  $Var[\bar{X} - \bar{Y}]$  and  $\sigma_{\bar{X} - \bar{Y}}$ 

### Proposition

The expected value of  $\overline{X} - \overline{Y}$  is  $\mu_1 - \mu_2$ , so  $\overline{X} - \overline{Y}$  is an unbiased estimator of  $\mu_1 - \mu_2$ . The standard deviation of  $\overline{X} - \overline{Y}$  is

$$\sigma_{\overline{X}-\overline{Y}} = \sqrt{\frac{\sigma_1^2}{m} + \frac{\sigma_2^2}{n}}$$

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### Normal distributions with known variances

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# Chapter 8: Confidence intervals

Assume further that the distributions of X and Y are normal and  $\sigma_1$ ,  $\sigma_2$  are known:

### Problem

(a) What is the distribution of

$$\frac{(\bar{X}-\bar{Y})-(\mu_1-\mu_2)}{\sqrt{\frac{\sigma_1^2}{m}+\frac{\sigma_2^2}{n}}}$$

(b) Compute

$$P\left[-1.96 \le \frac{(\bar{X} - \bar{Y}) - (\mu_1 - \mu_2)}{\sqrt{\frac{\sigma_1^2}{m} + \frac{\sigma_2^2}{n}}} \le 1.96\right]$$

(c) Construct a 95% CI for  $\mu_1 - \mu_2$  (in terms of  $\bar{x}$ ,  $\bar{y}$ , m, n,  $\sigma_1$ ,  $\sigma_2$ ).

When both population distributions are normal, standardizing  $\overline{X} - \overline{Y}$  gives a random variable Z with a standard normal distribution. Since the area under the z curve between  $-z_{\alpha/2}$  and  $z_{\alpha/2}$  is  $1 - \alpha$ , it follows that

$$P\left(-z_{\alpha/2} < \frac{\overline{X} - \overline{Y} - (\mu_1 - \mu_2)}{\sqrt{\frac{\sigma_1^2}{m} + \frac{\sigma_2^2}{n}}} < z_{\alpha/2}\right) = 1 - \alpha$$

Manipulation of the inequalities inside the parentheses to isolate  $\mu_1 - \mu_2$  yields the equivalent probability statement

$$P\left(\overline{X} - \overline{Y} - z_{\alpha/2}\sqrt{\frac{\sigma_1^2}{m} + \frac{\sigma_2^2}{n}} < \mu_1 - \mu_2 < \overline{X} - \overline{Y} + z_{\alpha/2}\sqrt{\frac{\sigma_1^2}{m} + \frac{\sigma_2^2}{n}}\right) = 1 - \alpha$$

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## Testing the difference between two population means

- Setting: independent normal random samples X<sub>1</sub>, X<sub>2</sub>,..., X<sub>m</sub> and Y<sub>1</sub>, Y<sub>2</sub>,..., Y<sub>n</sub> with known values of σ<sub>1</sub> and σ<sub>2</sub>. Constant Δ<sub>0</sub>.
- Null hypothesis:

$$H_0: \mu_1 - \mu_2 = \Delta_0$$

• Alternative hypothesis:

(a) 
$$H_a: \mu_1 - \mu_2 > \Delta_0$$
  
(b)  $H_a: \mu_1 - \mu_2 < \Delta_0$   
(c)  $H_a: \mu_1 - \mu_2 \neq \Delta_0$ 

• When  $\Delta = 0$ , the test (c) becomes

$$H_0: \mu_1 = \mu_2$$
$$H_a: \mu_1 \neq \mu_2$$

### Problem

Assume that we want to test the null hypothesis  $H_0: \mu_1 - \mu_2 = \Delta_0$ against each of the following alternative hypothesis

(a) 
$$H_a: \mu_1 - \mu_2 > \Delta_0$$
  
(b)  $H_a: \mu_1 - \mu_2 < \Delta_0$   
(c)  $H_a: \mu_1 - \mu_2 \neq \Delta_0$ 

by using the test statistic:

$$z = rac{(ar{x} - ar{y}) - (\mu_1 - \mu_2)}{\sqrt{rac{\sigma_1^2}{m} + rac{\sigma_2^2}{n}}}.$$

What is the rejection region in each case (a), (b) and (c)?

#### Proposition

Null hypothesis:  $H_0: \mu_1 - \mu_2 = \Delta_0$ Test statistic value:  $z = \frac{\overline{x} - \overline{y} - \Delta_0}{\sqrt{\frac{\sigma_1^2}{m} + \frac{\sigma_2^2}{n}}}$ 

Alternative Hypothesis

#### Rejection Region for Level a Test

- $\begin{array}{l} H_{\rm a}: \, \mu_1 \mu_2 > \Delta_0 \\ H_{\rm a}: \, \mu_1 \mu_2 < \Delta_0 \\ H_{\rm a}: \, \mu_1 \mu_2 \neq \Delta_0 \end{array}$
- $z \ge z_{\alpha} \text{ (upper-tailed test)}$  $z \le -z_{\alpha} \text{ (lower-tailed test)}$  $either <math>z \ge z_{\alpha/2} \text{ or } z \le -z_{\alpha/2} \text{ (two-tailed test)}$

Each student in a class of 21 responded to a questionnaire that requested their GPA and the number of hours each week that they studied. For those who studied less than 10 h/week the GPAs were

2.80, 3.40, 4.00, 3.60, 2.00, 3.00, 3.47, 2.80, 2.60, 2.00

and for those who studied at least 10 h/week the GPAs were

3.00, 3.00, 2.20, 2.40, 4.00, 2.96, 3.41, 3.27, 3.80, 3.10, 2.50

Assume that the distribution of GPA for each group is normal and both distributions have standard deviation  $\sigma_1 = \sigma_2 = 0.6$ . Treating the two samples as random, is there evidence that true average GPA differs for the two study times? Carry out a test of significance at level .05.

## Solution

- The parameter of interest is µ<sub>1</sub> − µ<sub>2</sub>, the difference between true mean GPA for the < 10 (conceptual) population and true mean GPA for the ≥10 population.</li>
- 2. The null hypothesis is  $H_0: \mu_1 \mu_2 = 0$ .
- 3. The alternative hypothesis is H<sub>a</sub>: µ<sub>1</sub> − µ<sub>2</sub> ≠ 0; if H<sub>a</sub> is true then µ<sub>1</sub> and µ<sub>2</sub> are different. Although it would seem unlikely that µ<sub>1</sub> − µ<sub>2</sub> > 0 (those with low study hours have higher mean GPA) we will allow it as a possibility and do a two-tailed test.
- 4. With  $\Delta_0 = 0$ , the test statistic value is

$$z = \frac{\overline{x} - \overline{y}}{\sqrt{\frac{\sigma_1^2}{m} + \frac{\sigma_2^2}{n}}}$$

5. The inequality in  $H_a$  implies that the test is two-tailed. For  $\alpha = .05$ ,  $\alpha/2 = .025$  and  $z_{\alpha/2} = z_{.025} = 1.96$ .  $H_0$  will be rejected if  $z \ge 1.96$  or  $z \le -1.96$ .

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## Solution

6. Substituting m = 10,  $\bar{x} = 2.97$ ,  $\sigma_1^2 = .36$ , n = 11,  $\bar{y} = 3.06$ , and  $\sigma_2^2 = .36$  into the formula for z yields

$$z = \frac{2.97 - 3.06}{\sqrt{\frac{.36}{10} + \frac{.36}{11}}} = \frac{-.09}{.262} = -.34$$

That is, the value of  $\overline{x} - \overline{y}$  is only one-third of a standard deviation below what would be expected when  $H_0$  is true.

 Because the value of z is not even close to the rejection region, there is no reason to reject the null hypothesis. This test shows no evidence of any relationship between study hours and GPA.

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## Large-sample tests/confidence intervals

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# Principles

• Central Limit Theorem:  $\bar{X}$  and  $\bar{Y}$  are approximately normal when  $n > 30 \rightarrow$  so is  $\bar{X} - \bar{Y}$ . Thus

$$\frac{(\bar{X}-\bar{Y})-(\mu_1-\mu_2)}{\sqrt{\frac{\sigma_1^2}{m}+\frac{\sigma_2^2}{n}}}$$

is approximately standard normal

- When *n* is sufficiently large  $S_1 \approx \sigma_1$  and  $S_2 \approx \sigma_2$
- Conclusion:

$$\frac{(\bar{X} - \bar{Y}) - (\mu_1 - \mu_2)}{\sqrt{\frac{S_1^2}{m} + \frac{S_2^2}{n}}}$$

is approximately standard normal when *n* is sufficiently large If m, n > 40, we can ignore the normal assumption and replace  $\sigma$  by *S* 

#### Proposition

Use of the test statistic value

$$z = \frac{\overline{x} - \overline{y} - \Delta_0}{\sqrt{\frac{s_1^2}{m} + \frac{s_2^2}{n}}}$$

along with the previously stated upper-, lower-, and two-tailed rejection regions based on z critical values gives large-sample tests whose significance levels are approximately  $\alpha$ . These tests are usually appropriate if both m > 40 and n > 40. A P-value is computed exactly as it was for our earlier z tests.

### Proposition

Provided that m and n are both large, a CI for  $\mu_1 - \mu_2$  with a confidence level of approximately  $100(1 - \alpha)\%$  is

$$ar{x} - ar{y} \pm z_{lpha/2} \sqrt{rac{s_1^2}{m} + rac{s_2^2}{n}}$$

where -gives the lower limit and + the upper limit of the interval. An upper or lower confidence bound can also be calculated by retaining the appropriate sign and replacing  $z_{\alpha/2}$  by  $z_{\alpha}$ .

#### Example

Let  $\mu_1$  and  $\mu_2$  denote true average tread lives for two competing brands of size P205/65R15 radial tires.

(a) Test

 $H_0: \mu_1 = \mu_2$  $H_a: \mu_1 \neq \mu_2$ 

at level 0.05 using the following data: m = 45,  $\bar{x} = 42,500$ ,  $s_1 = 2200$ , n = 45,  $\bar{y} = 40,400$ , and  $s_2 = 1900$ . (b) Construct a 95% CI for  $\mu_1 - \mu_2$ .