## MATH 450: Mathematical statistics

Oct 13th, 2020
Lecture 13: Confidence intervals

## Countdown to midterm: 16 days

| Week 2 | Chapter 6: Statistics and Sampling Distributions |
| :---: | :---: |
| Week 4 | Chapter 7: Point Estimation |
| Week 7 | Chapter 8: Confidence Intervals |
| Week 10 | Chapter 9: Test of Hypothesis |
| Week 11 | Chapter 10: Two-sample inference |
| Week 13 | Regression |

## Chapter 8: Confidence intervals

## Overview

8.1 Basic properties of confidence intervals (Cls)

- Interpreting Cls
- General principles to derive Cl
8.2 Large-sample confidence intervals for a population mean
- Using the Central Limit Theorem to derive Cls
8.3 Intervals based on normal distribution
- Using Student's t-distribution
8.4 Cls for standard deviation
- Let $X_{1}, X_{2}, \ldots, X_{n}$ be a random sample from a distribution $f(x, \theta)$
- In Chapter 7, we learnt methods to construct an estimate $\hat{\theta}$ of $\theta$
- Goal: we want to indicate the degree of uncertainty associated with this random prediction
- One way to do so is to construct a confidence interval $[\hat{\theta}-a, \hat{\theta}+b]$ such that

$$
P[\theta \in[\hat{\theta}-a, \hat{\theta}+b]]=95 \%
$$

## Confidence interval



## Principles for deriving Cls

If $X_{1}, X_{2}, \ldots, X_{n}$ is a random sample from a distribution $f(x, \theta)$, then

- Find a random variable $Y=h\left(X_{1}, X_{2}, \ldots, X_{n} ; \theta\right)$ such that the probability distribution of $Y$ does not depend on $\theta$ or on any other unknown parameters.
- Find constants $a, b$ such that

$$
P\left[a<h\left(X_{1}, X_{2}, \ldots, X_{n} ; \theta\right)<b\right]=0.95
$$

- Manipulate these inequalities to isolate $\theta$

$$
P\left[\ell\left(X_{1}, X_{2}, \ldots, X_{n}\right)<\theta<u\left(X_{1}, X_{2}, \ldots, X_{n}\right)\right]=0.95
$$

## Confidence interval: Example 1

## Problem

Suppose the sediment density $(\mathrm{g} / \mathrm{cm})$ of a randomly selected specimen from a certain region is normally distributed with mean $\mu$ (unknown) and standard deviation $\sigma=0.85$. A random sample of $n=25$ specimens is selected with sample average $\bar{X}$.

- (Step 1) What is the distribution of

$$
\frac{\bar{X}-\mu}{\sigma / \sqrt{n}}
$$

- (Step 2) Find a number c such that

$$
P\left[-c<\frac{\bar{X}-\mu}{\sigma / \sqrt{n}}<c\right]=0.95
$$

| $z$ |  |  | .00 | .01 | .02 | .03 | .04 | .05 | .06 | .07 |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| .08 | .09 |  |  |  |  |  |  |  |  |  |
| 0.0 | .5000 | .5040 | .5080 | .5120 | .5160 | .5199 | .5239 | .5279 | .5319 | .5359 |
| 0.1 | .5398 | .5438 | .5478 | .5517 | .5557 | .5596 | .5636 | .5675 | .5714 | .5753 |
| 0.2 | .5793 | .5832 | .5871 | .5910 | .5948 | .5987 | .6026 | .6064 | .6103 | .6141 |
| 0.3 | .6179 | .6217 | .6255 | .6293 | .6331 | .6368 | .6406 | .6443 | .6480 | .6517 |
| 0.4 | .6554 | .6591 | .6628 | .6664 | .6700 | .6736 | .6772 | .6808 | .6844 | .6879 |
| 0.5 | .6915 | .6950 | .6985 | .7019 | .7054 | .7088 | .7123 | .7157 | .7190 | .7224 |
| 0.6 | .7257 | .7291 | .7324 | .7357 | .7389 | .7422 | .7454 | .7486 | .7517 | .7549 |
| 0.7 | .7580 | .7611 | .7642 | .7673 | .7704 | .7734 | .7764 | .7794 | .7823 | .7852 |
| 0.8 | .7881 | .7910 | .7939 | .7967 | .7995 | .8023 | .8051 | .8078 | .8106 | .8133 |
| 0.9 | .8159 | .8186 | .8212 | .8238 | .8264 | .8289 | .8315 | .8340 | .8365 | .8389 |
| 1.0 | .8413 | .8438 | .8461 | .8485 | .8508 | .8531 | .8554 | .8577 | .8599 | .8621 |
| 1.1 | .8643 | .8665 | .8686 | .8708 | .8729 | .8749 | .8770 | .8790 | .8810 | .8830 |
| 1.2 | .8849 | .8869 | .8888 | .8907 | .8925 | .8944 | .8962 | .8980 | .8997 | .9015 |
| 1.3 | .032 | .9049 | .9066 | .9082 | .9099 | .9115 | .9131 | .9147 | .9162 | .9177 |
| 1.4 | .9192 | .9207 | .9222 | .9236 | .9251 | .9265 | .9278 | .9292 | .9306 | .9319 |
| 1.5 | .9332 | .9345 | .9357 | .9370 | .9382 | .9394 | .9406 | .9418 | .9429 | .9441 |
| 1.6 | .9452 | .9463 | .9474 | .9484 | .9495 | .9505 | .9515 | .9525 | .9535 | .9545 |
| 1.7 | .9554 | .9564 | .9573 | .9582 | .9591 | .9599 | .9608 | .9616 | .9625 | .9633 |
| 1.8 | .9641 | .9649 | .9656 | .9664 | .9671 | .9678 | .9686 | .9693 | .9699 | .9706 |
| 1.9 | .9713 | .9719 | .9726 | .9732 | .9738 | .9744 | .9750 | .9756 | .9761 | .9767 |
| 2.0 | .9772 | .9778 | .9783 | .9788 | .9793 | .9798 | .9803 | .9808 | .9812 | .9817 |
| 2.1 | .9821 | .9826 | .9830 | .9834 | .9838 | .9842 | .9846 | .9850 | .9854 | .9857 |
| 2.2 | .9861 | .9864 | .9868 | .871 | .9875 | .9878 | .9881 | .9884 | .9887 | .9890 |
| 2.3 | .893 | .9896 | .9898 | .9901 | .9904 | .9906 | .9909 | .9911 | .9913 | .9916 |
| 2.4 | .9918 | .9920 | .9922 | .9925 | .9927 | .9929 | .9931 | .9932 | .9934 | .9936 |
| 2.5 | .9938 | .9940 | .9941 | .9943 | .9945 | .9946 | .9948 | .9949 | .9951 | .9952 |
| 2.6 | .9953 | .9955 | .9956 | .9957 | .9959 | .9960 | .9961 | .9962 | .9963 | .9964 |
| 2.7 | .9965 | .9966 | .9967 | .9968 | .9969 | .9970 | .9971 | .9972 | .9973 | .9974 |
| 2.8 | .9974 | .9975 | .9976 | .9977 | .9977 | .9978 | .9979 | .9979 | .9980 | .9981 |
| 2.9 | .9981 | .9982 | .9982 | .9983 | .9984 | .9984 | .9985 | .9985 | .9986 | .9986 |
| 3.0 | .9987 | .9987 | .9987 | .9988 | .9988 | .9989 | .9989 | .9989 | .9990 | .9990 |
| 3.1 | .999 | .9991 | .9991 | .991 | .9992 | .9992 | .9992 | .9992 | .9993 | .9993 |
| 3.2 | .9993 | .9993 | .9994 | .9994 | .9994 | .9994 | .9994 | .9995 | .9995 | .9995 |
| 3.3 | .9995 | .9995 | .9995 | .9996 | .9996 | .9996 | .9996 | .9996 | .9996 | .9997 |

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- Assumptions:
- Normal distribution
- $\sigma$ is known
- $95 \%$ confidence interval

If after observing $X_{1}=x_{1}, X_{2}=x_{2}, \ldots, X_{n}=x_{n}$, we compute the observed sample mean $\bar{x}$. Then

$$
\left(\bar{x}-1.96 \frac{\sigma}{\sqrt{n}}, \bar{x}+1.96 \frac{\sigma}{\sqrt{n}}\right)
$$

is a $95 \%$ confidence interval of $\mu$

NOTATION
$z_{\alpha}$ will denote the value on the measurement axis for which $\alpha$ of the area under the $z$ curve lies to the right of $z_{\alpha}$. (See Figure 4.19.)

For example, $z_{.10}$ captures upper-tail area .10 and $z_{.01}$ captures upper-tail area 01 .


Figure $4.19 z_{\alpha}$ notation illustrated
Since $\alpha$ of the area under the standard normal curve lies to the right of $z_{\alpha}, 1-\alpha$ of the area lies to the left of $z_{\alpha}$. Thus $z_{\alpha}$ is the $100(1-\alpha)$ th percentile of the standard normal distribution. By symmetry the area under the standard normal curve to the left of $-z_{\alpha}$ is also $\alpha$. The $z_{\alpha}$ 's are usually referred to as $z$ critical values. Table 4.1 lists the most useful standard normal percentiles and $z_{\alpha}$ values.

## $100(1-\alpha) \%$ confidence interval



Figure 8.4 $P\left(-z_{\alpha / 2} \leq Z \leq z_{\alpha / 2}\right)=1-\alpha$

## $100(1-\alpha) \%$ confidence interval

A $100(1-\alpha) \%$ confidence interval for the mean $\mu$ of a normal population when the value of $\sigma$ is known is given by

$$
\begin{equation*}
\left(\bar{x}-z_{\alpha / 2} \cdot \frac{\sigma}{\sqrt{n}}, \bar{x}+z_{\alpha / 2} \cdot \frac{\sigma}{\sqrt{n}}\right) \tag{8.5}
\end{equation*}
$$

or, equivalently, by $\bar{x} \pm z_{\alpha / 2} \cdot \sigma / \sqrt{n}$.

## Interpreting confidence intervals



95\% confidence interval: If we repeat the experiment many times, the interval contains $\mu$ about $95 \%$ of the time

## Interpreting confidence intervals

- Writing

$$
P[\mu \in(\bar{X}-1.7, \bar{X}+1.7)]=95 \%
$$

is okay.

- If $\bar{x}=2.7$, writing

$$
P[\mu \in(1,4.4)]=95 \%
$$

is NOT okay.

- Saying $\mu \in(1,4.4)$ with confidence level $95 \%$ is okay.
- Saying "if we repeat the experiment many times, the interval contains $\mu$ about $95 \%$ of the time" is perfect.


## Example 2

## Example

Assume that the helium porosity (in percentage) of coal samples taken from any particular seam is normally distributed with true standard deviation $\sigma=.75$.

- Compute a $95 \% \mathrm{Cl}$ for the true average porosity of a certain seam if the average porosity for 20 specimens from the seam was 4.85 .
- How large a sample size is necessary if the width of the $95 \%$ interval is to be .40 ?


## One-sided Cls (Confidence bounds)

## Example 1b: One-sided confidence interval

## Problem

Suppose the sediment density $(\mathrm{g} / \mathrm{cm})$ of a randomly selected specimen from a certain region is normally distributed with mean $\mu$ (unknown) and standard deviation $\sigma=0.85$. A random sample of $n=25$ specimens is selected with sample average $\bar{X}$.

- (Step 1) What is the distribution of

$$
\frac{\bar{x}-\mu}{\sigma / \sqrt{n}}
$$

- (Step 2) Find a number b such that

$$
P\left[\frac{\bar{X}-\mu}{\sigma / \sqrt{n}}<b\right]=0.95
$$

## Cls vs. one-sided Cls

Cls:

- 100(1- $\alpha$ )\% confidence

$$
\left(\bar{x}-z_{\alpha / 2} \frac{s}{\sqrt{n}}, \bar{x}+z_{\alpha / 2} \frac{s}{\sqrt{n}}\right)
$$

- $95 \%$ confidence

$$
\left(\bar{x}-1.96 \frac{s}{\sqrt{n}}, \bar{x}+1.96 \frac{s}{\sqrt{n}}\right)
$$

One-sided Cls:

- 100(1- $\alpha$ ) \% confidence

$$
\left(-\infty, \bar{x}+z_{\alpha} \frac{s}{\sqrt{n}}\right)
$$

- $95 \%$ confidence

$$
\left(-\infty, \bar{x}+1.64 \frac{s}{\sqrt{n}}\right)
$$

