## MATH 450: Mathematical statistics

$$
\text { Oct 22nd, } 2020
$$

Lecture 16: Confidence intervals for standard deviation

## Countdown to midterm: 7 days

| Week 2 | Chapter 6: Statistics and Sampling Distributions |
| :---: | :---: |
| Week 4 | Chapter 7: Point Estimation |
| Week 7 | Chapter 8: Confidence Intervals |
| Week 10 | Chapter 9: Test of Hypothesis |
| Week 11 | Chapter 10: Two-sample inference |
| Week 13 | Regression |

## Overview

8.1 Basic properties of confidence intervals (Cls)

- Interpreting Cls
- General principles to derive Cl
8.2 Large-sample confidence intervals for a population mean
- Using the Central Limit Theorem to derive Cls
8.3 Intervals based on normal distribution
- Using Student's t-distribution
8.4 Cls for standard deviation


## Confidence intervals for a population mean

- Section 8.1: Normal distribution with known $\sigma$
- Normal distribution
- $\sigma$ is known
- Section 8.2: Large-sample confidence intervals
- Normal distribution
$\rightarrow$ use Central Limit Theorem $\rightarrow$ needs $n>30$
- $\sigma$ is known
$\rightarrow$ replace $\sigma$ by $s \rightarrow$ needs $n>40$
- Section 8.3: Intervals based on normal distributions
- Normal distribution
- $\sigma$ is known
$\rightarrow$ Introducing $t$-distribution


## 8.1: Normal distribution with known $\sigma$

- Assumptions:
- Normal distribution
- $\sigma$ is known
- $95 \%$ confidence interval

If after observing $X_{1}=x_{1}, X_{2}=x_{2}, \ldots, X_{n}=x_{n}$, we compute the observed sample mean $\bar{x}$. Then

$$
\left(\bar{x}-1.96 \frac{\sigma}{\sqrt{n}}, \bar{x}+1.96 \frac{\sigma}{\sqrt{n}}\right)
$$

is a $95 \%$ confidence interval of $\mu$
$z_{\alpha}$ will denote the value on the measurement axis for which $\alpha$ of the area under the $z$ curve lies to the right of $z_{\alpha}$. (See Figure 4.19.)

For example, $z_{.10}$ captures upper-tail area .10 and $z_{.01}$ captures upper-tail area 01 .


Figure $4.19 z_{\alpha}$ notation illustrated
Since $\alpha$ of the area under the standard normal curve lies to the right of $z_{\alpha}, 1-\alpha$ of the area lies to the left of $z_{\alpha}$. Thus $z_{\alpha}$ is the $100(1-\alpha)$ th percentile of the standard normal distribution. By symmetry the area under the standard normal curve to the left of $-z_{\alpha}$ is also $\alpha$. The $z_{\alpha}$ 's are usually referred to as $z$ critical values. Table 4.1 lists the most useful standard normal percentiles and $z_{\alpha}$ values.

## $100(1-\alpha) \%$ confidence interval

A $100(1-\alpha) \%$ confidence interval for the mean $\mu$ of a normal population when the value of $\sigma$ is known is given by

$$
\begin{equation*}
\left(\bar{x}-z_{\alpha / 2} \cdot \frac{\sigma}{\sqrt{n}}, \bar{x}+z_{\alpha / 2} \cdot \frac{\sigma}{\sqrt{n}}\right) \tag{8.5}
\end{equation*}
$$

or, equivalently, by $\bar{x} \pm z_{\alpha / 2} \cdot \sigma / \sqrt{n}$.

## 8.2: Large-sample Cls of the population mean

## Principles

- Central Limit Theorem

$$
\frac{\bar{x}-\mu}{\sigma / \sqrt{n}}
$$

is approximately normal when $n>30$

- Moreover, when $n$ is sufficiently large $s \approx \sigma$
- Conclusion:

$$
\frac{\bar{X}-\mu}{s / \sqrt{n}}
$$

is approximately normal when $n$ is sufficiently large
If $n>40$, we can ignore the normal assumption and replace $\sigma$ by $s$

## $100(1-\alpha) \%$ confidence interval

If after observing $X_{1}=x_{1}, X_{2}=x_{2}, \ldots, X_{n}=x_{n}(n>40)$, we compute the observed sample mean $\bar{x}$ and sample standard deviation s. Then

$$
\left(\bar{x}-z_{\alpha / 2} \frac{s}{\sqrt{n}}, \bar{x}+z_{\alpha / 2} \frac{s}{\sqrt{n}}\right)
$$

is a $95 \%$ confidence interval of $\mu$

## One-sided Cls

## A large-sample upper confidence bound for $\mu$ is

$$
\mu<\bar{x}+z_{\alpha} \cdot \frac{s}{\sqrt{n}}
$$

and a large-sample lower confidence bound for $\mu$ is

$$
\mu>\bar{x}-z_{\alpha} \cdot \frac{s}{\sqrt{n}}
$$

## 8.3: Intervals based on normal distributions

## Assumptions

- the population of interest is normal (i.e., $X_{1}, \ldots, X_{n}$ constitutes a random sample from a normal distribution $\left.\mathcal{N}\left(\mu, \sigma^{2}\right)\right)$.
- $\sigma$ is unknown
$\rightarrow$ we want to consider cases when $n$ is small.


## $t$ distributions

When $\bar{X}$ is the mean of a random sample of size n from a normal distribution with mean $\mu$, the rv

$$
\frac{\bar{x}-\mu}{S / \sqrt{n}}
$$

has the $t$ distribution with $n-1$ degree of freedom (df).

## Confidence intervals

Let $\bar{x}$ and $s$ be the sample mean and sample standard deviation computed from the results of a random sample from a normal population with mean $\mu$. Then a $\mathbf{1 0 0}(1-\alpha) \%$ confidence interval for $\boldsymbol{\mu}$, the one-sample $\boldsymbol{t}$ CI, is

$$
\begin{equation*}
\left(\bar{x}-t_{\alpha / 2, n-1} \cdot \frac{s}{\sqrt{n}}, \bar{x}+t_{\alpha / 2, n-1} \cdot \frac{s}{\sqrt{n}}\right) \tag{8.15}
\end{equation*}
$$

or, more compactly, $\bar{x} \pm t_{\alpha / 2, n-1} \cdot s / \sqrt{n}$.
An upper confidence bound for $\boldsymbol{\mu}$ is

$$
\bar{x}+t_{\alpha, n-1} \cdot \frac{s}{\sqrt{n}}
$$

and replacing + by - in this latter expression gives a lower confidence bound for $\boldsymbol{\mu}$; both have confidence level $100(1-\alpha) \%$.

Let $t_{\alpha, v}=$ the number on the measurement axis for which the area under the $t$ curve with $v$ df to the right of $t_{\alpha, v}$, is $\alpha ; t_{\alpha, v}$ is called a $t$ critical value.


Figure 8.7 A pictorial definition of $t_{\alpha, \nu}$

Table A. 5 Critical Values for $t$ Distributions

$\alpha$

| $\nu$ | . 10 | . 05 | . 025 | . 01 | . 005 | . 001 | . 0005 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 1 | 3.078 | 6.314 | 12.706 | 31.821 | 63.657 | 318.31 | 636.62 |
| 2 | 1.886 | 2.920 | 4.303 | 6.965 | 9.925 | 22.326 | 31.598 |
| 3 | 1.638 | 2.353 | 3.182 | 4.541 | 5.841 | 10.213 | 12.924 |
| 4 | 1.533 | 2.132 | 2.776 | 3.747 | 4.604 | 7.173 | 8.610 |
| 5 | 1.476 | 2.015 | 2.571 | 3.365 | 4.032 | 5.893 | 6.869 |
| 6 | 1.440 | 1.943 | 2.447 | 3.143 | 3.707 | 5.208 | 5.959 |
| 7 | 1.415 | 1.895 | 2.365 | 2.998 | 3.499 | 4.785 | 5.408 |
| 8 | 1.397 | 1.860 | 2.306 | 2.896 | 3.355 | 4.501 | 5.041 |
| 9 | 1.383 | 1.833 | 2.262 | 2.821 | 3.250 | 4.297 | 4.781 |
| 10 | 1.372 | 1.812 | 2.228 | 2.764 | 3.169 | 4.144 | 4.587 |
| 11 | 1.363 | 1.796 | 2.201 | 2.718 | 3.106 | 4.025 | 4.437 |
| 12 | 1.356 | 1.782 | 2.179 | 2.681 | 3.055 | 3.930 | 4.318 |
| 13 | 1.350 | 1.771 | 2.160 | 2.650 | 3.012 | 3.852 | 4.221 |
| 14 | 1.345 | 1.761 | 2.145 | 2.624 | 2.977 | 3.787 | 4.140 |
| 15 | 1.341 | 1.753 | 2.131 | 2.602 | 2.947 | 3.733 | 4.073 |
| 16 | 1.337 | 1.746 | 2.120 | 2.583 | 2.921 | 3.686 | 4.015 |
| 17 | 1.333 | 1.740 | 2.110 | 2.567 | 2.898 | 3.646 | 3.965 |
| 18 | 1.330 | 1.734 | 2.101 | 2.552 | 2.878 | 3.610 | 3.922 |
| 19 | 1.328 | 1.729 | 2.093 | 2.539 | 2.861 | 3.579 | 3.883 |
| 20 | 1.325 | 1.725 | 2.086 | 2.528 | 2.845 | 3.552 | 3.850 |
| 21 | 1.323 | 1.721 | 2.080 | 2.518 | 2.831 | 3.527 | 3.819 |
| 22 | 1.321 | 1.717 | 2.074 | 2.508 | 2.819 | 3.505 | 3.792 |
| 23 | 1.319 | 1.714 | 2.069 | 2.500 | 2.807 | 3.485 | 3.767 |
| 24 | 1.318 | 1.711 | 2.064 | 2.492 | 2.797 | 3.467 | 3.745 |

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## Prediction intervals

## Settings

- We have available a random sample $X_{1}, X_{2}, \ldots, X_{n}$ from a normal population distribution
- We wish to predict the value of $X_{n+1}$, a single future observation.

This is a much more difficult problem than the problem of estimating $\mu$

- When $n \rightarrow \infty, \bar{X} \rightarrow \mu$
- Even when we know $\mu, X_{n+1}$ is still random

Let $X_{1}, X_{2}, \ldots, X_{n}$ be a sample from a normal population distribution $\mathcal{N}(\mu, \sigma)$ and $X_{n+1}$ be an independent sample from the same distribution.

- Compute $E\left[\bar{X}-X_{n+1}\right]$ in terms of $\mu, \sigma, n$
- Compute $\operatorname{Var}\left[\bar{X}-X_{n+1}\right]$ in terms of $\mu, \sigma, n$
- What is the distribution of $\bar{X}-X_{n+1}$ ?


## Principle

If $\sigma$ is known

$$
\frac{\bar{X}-X_{n+1}}{\sigma \sqrt{1+\frac{1}{n}}}
$$

follows the standard normal distribution $\mathcal{N}(0,1)$.

## Principle

$$
T=\frac{X-X_{n+1}}{S \sqrt{1+\frac{1}{n}}} \sim t \text { distribution with } n-1 \mathrm{df}
$$

## Prediction intervals

A prediction interval (PI) for a single observation to be selected from a normal population distribution is

$$
\begin{equation*}
\bar{x} \pm t_{\alpha / 2, n-1} \cdot s \sqrt{1+\frac{1}{n}} \tag{8.16}
\end{equation*}
$$

The prediction level is $100(1-\alpha) \%$.

## Example 4b

## Example

Let $X$ be the amount of butterfat in pounds produced by a typical cow during a 305-day milk production period between her first and second calves. Assume that the distribution of X is $N\left(\mu, \sigma^{2}\right)$. To estimate $\mu$, a farmer measured the butterfat production for $\mathrm{n}=20$ cows and obtained the following data:

| 481 | 537 | 513 | 583 | 453 | 510 | 570 | 500 | 457 | 555 |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- |
| 618 | 327 | 350 | 643 | 499 | 421 | 505 | 637 | 599 | 392 |

Construct a $90 \%$ prediction interval for $\mu$.

## Section 6.4: Distributions based on a normal random sample

- The Chi-squared distribution
- The $t$ distribution
- The F Distribution


## Chi-squared distribution

The pdf of a Chi-squared distribution with degree of freedom $\nu$, denoted by $\chi_{\nu}^{2}$, is

$$
f(x)=\left\{\begin{array}{cc}
\frac{1}{2^{1 / 2} \Gamma(v / 2)} x^{(v / 2)-1} e^{-x / 2} & x>0 \\
0 & x \leq 0
\end{array}\right.
$$



## Why is Chi-squared useful?

## Proposition

If $Z$ has standard normal distribution $\mathcal{Z}(0,1)$ and $X=Z^{2}$, then $X$ has Chi-squared distribution with 1 degree of freedom, i.e. $X \sim \chi_{1}^{2}$ distribution.

## Proposition

If $X_{1} \sim \chi_{\nu_{1}}^{2}, X_{2} \sim \chi_{\nu_{2}}^{2}$ and they are independent, then

$$
X_{1}+X_{2} \sim \chi_{\nu_{1}+\nu_{2}}^{2}
$$

## Why is Chi-squared useful?

## Proposition

If $Z_{1}, Z_{2}, \ldots, Z_{n}$ are independent and each has the standard normal distribution, then

$$
Z_{1}^{2}+Z_{2}^{2}+\ldots+Z_{n}^{2} \sim \chi_{n}^{2}
$$

## Why is Chi-squared useful?

If $X_{1}, X_{2}, \ldots, X_{n}$ is a random sample from the normal distribution $\mathcal{N}\left(\mu, \sigma^{2}\right)$, then

$$
(n-1) \frac{S^{2}}{\sigma^{2}} \sim \chi_{n-1}^{2}
$$

Let $Z$ be a standard normal rv and let $X$ be a $\chi_{\nu}^{2}$ rv independent of $Z$. Then the $t$ distribution with degrees of freedom $\nu$ is defined to be the distribution of the ratio

$$
T=\frac{Z}{\sqrt{X / \nu}}
$$

When $\bar{X}$ is the mean of a random sample of size n from a normal distribution with mean $\mu$, the rv

$$
\frac{\bar{X}-\mu}{S / \sqrt{n}}
$$

has the $t$ distribution with $n-1$ degree of freedom (df). Hint:

$$
T=\frac{Z}{\sqrt{X / \nu}} \quad(n-1) \frac{S^{2}}{\sigma^{2}} \sim \chi_{n-1}^{2}
$$

and

$$
\frac{\bar{X}-\mu}{S / \sqrt{n}}=\frac{\bar{X}-\mu}{\sigma / \sqrt{n}} \cdot \frac{1}{\sqrt{(n-1) \frac{S^{2}}{\sigma^{2}} /(n-1)}} .
$$

## Cls for variance and standard deviation

## Why is Chi-squared useful?

If $X_{1}, X_{2}, \ldots, X_{n}$ is a random sample from the normal distribution $\mathcal{N}\left(\mu, \sigma^{2}\right)$, then

$$
(n-1) \frac{S^{2}}{\sigma^{2}} \sim \chi_{n-1}^{2}
$$

## Important: Chi-squared distribution are not symmetric



## Cls for standard deviation

We have

$$
P\left(\chi_{1-\alpha / 2, n-1}^{2}<\frac{(n-1) S^{2}}{\sigma^{2}}<\chi_{\alpha / 2, n-1}^{2}\right)=1-\alpha
$$

Play around with these inequalities:

$$
\frac{(n-1) S^{2}}{\chi_{\alpha / 2, n-1}^{2}}<\sigma^{2}<\frac{(n-1) S^{2}}{\chi_{1-\alpha / 2, n-1}^{2}}
$$

## Cls for standard deviation

A $\mathbf{1 0 0}(1-\alpha) \%$ confidence interval for the variance $\boldsymbol{\sigma}^{\mathbf{2}}$ of a normal population has lower limit

$$
(n-1) s^{2} / \chi_{\alpha / 2, n-1}^{2}
$$

and upper limit

$$
(n-1) s^{2} / \chi_{1-\alpha / 2, n-1}^{2}
$$

A confidence interval for $\boldsymbol{\sigma}$ has lower and upper limits that are the square roots of the corresponding limits in the interval for $\boldsymbol{\sigma}^{2}$.

## Practice problems

## Example 1

## Problem

Here are the alcohol percentages for a sample of 16 beers:

| 4.68 | 4.13 | 4.80 | 4.63 | 5.08 | 5.79 | 6.29 | 6.79 |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- |
| 4.93 | 4.25 | 5.70 | 4.74 | 5.88 | 6.77 | 6.04 | 4.95 |

(a) Assume the distribution is normal, construct the 95\% confidence interval for the population mean.

Table A. 5 Critical Values for $t$ Distributions

$\alpha$

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| 3 | 1.638 | 2.353 | 3.182 | 4.541 | 5.841 | 10.213 | 12.924 |
| 4 | 1.533 | 2.132 | 2.776 | 3.747 | 4.604 | 7.173 | 8.610 |
| 5 | 1.476 | 2.015 | 2.571 | 3.365 | 4.032 | 5.893 | 6.869 |
| 6 | 1.440 | 1.943 | 2.447 | 3.143 | 3.707 | 5.208 | 5.959 |
| 7 | 1.415 | 1.895 | 2.365 | 2.998 | 3.499 | 4.785 | 5.408 |
| 8 | 1.397 | 1.860 | 2.306 | 2.896 | 3.355 | 4.501 | 5.041 |
| 9 | 1.383 | 1.833 | 2.262 | 2.821 | 3.250 | 4.297 | 4.781 |
| 10 | 1.372 | 1.812 | 2.228 | 2.764 | 3.169 | 4.144 | 4.587 |
| 11 | 1.363 | 1.796 | 2.201 | 2.718 | 3.106 | 4.025 | 4.437 |
| 12 | 1.356 | 1.782 | 2.179 | 2.681 | 3.055 | 3.930 | 4.318 |
| 13 | 1.350 | 1.771 | 2.160 | 2.650 | 3.012 | 3.852 | 4.221 |
| 14 | 1.345 | 1.761 | 2.145 | 2.624 | 2.977 | 3.787 | 4.140 |
| 15 | 1.341 | 1.753 | 2.131 | 2.602 | 2.947 | 3.733 | 4.073 |
| 16 | 1.337 | 1.746 | 2.120 | 2.583 | 2.921 | 3.686 | 4.015 |
| 17 | 1.333 | 1.740 | 2.110 | 2.567 | 2.898 | 3.646 | 3.965 |
| 18 | 1.330 | 1.734 | 2.101 | 2.552 | 2.878 | 3.610 | 3.922 |
| 19 | 1.328 | 1.729 | 2.093 | 2.539 | 2.861 | 3.579 | 3.883 |
| 20 | 1.325 | 1.725 | 2.086 | 2.528 | 2.845 | 3.552 | 3.850 |
| 21 | 1.323 | 1.721 | 2.080 | 2.518 | 2.831 | 3.527 | 3.819 |
| 22 | 1.321 | 1.717 | 2.074 | 2.508 | 2.819 | 3.505 | 3.792 |
| 23 | 1.319 | 1.714 | 2.069 | 2.500 | 2.807 | 3.485 | 3.767 |
| 24 | 1.318 | 1.711 | 2.064 | 2.492 | 2.797 | 3.467 | 3.745 |

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## Example 1b

## Problem

Here are the alcohol percentages for a sample of 16 beers:

| 4.68 | 4.13 | 4.80 | 4.63 | 5.08 | 5.79 | 6.29 | 6.79 |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- |
| 4.93 | 4.25 | 5.70 | 4.74 | 5.88 | 6.77 | 6.04 | 4.95 |

(b) Assume the distribution is normal, construct the 95\% lower confidence bound for the population mean.

## Example 1c

## Problem

Here are the alcohol percentages for a sample of 16 beers:

| 4.68 | 4.13 | 4.80 | 4.63 | 5.08 | 5.79 | 6.29 | 6.79 |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- |
| 4.93 | 4.25 | 5.70 | 4.74 | 5.88 | 6.77 | 6.04 | 4.95 |

(b) Assume that another beer is sampled from the same distribution, construct the 95\% prediction interval for the alcohol percentages of that beer.

## Example 2

## Problem

Suppose that against a certain opponent, the number of points a basketball team scores is normally distributed with unknown mean $\mu$ and unknown variance $\sigma^{2}$. Suppose that over the course of the last 10 games, the team scored the following points:

$$
59,62,59,74,70,61,62,66,62,75
$$

- Construct a 95\% confidence interval for $\mu$.
- Now suppose that you learn that $\sigma^{2}=25$. Construct a $95 \%$ confidence interval for $\mu$. How does this compare to the interval in (a)?


## Example 3

## Problem

A study of the ability of individuals to walk in a straight line reported the accompanying data on cadence (strides per second) for a sample of $n=20$ randomly selected healthy men. Assuming that the distribution is normal:

- Calculate a 95\% confidence interval for population mean cadence
- Calculate and interpret a 95\% prediction interval for the cadence of a single individual randomly selected from this population.

