

MATH 450: Mathematical statistics

November 10th, 2020

Lecture 19: Tests about a population mean

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| Week 2 | ● | Chapter 6: Statistics and Sampling Distributions |
| Week 4 | ● | Chapter 7: Point Estimation |
| Week 7 | ● | Chapter 8: Confidence Intervals |
| Week 10 | ● | Chapter 9: Tests of Hypotheses |
| Week 12 | ● | Chapter 10: Two-sample testing |
| Week 14 | ● | Regression |

9.1 Hypotheses and test procedures

- test procedures
- errors in hypothesis testing
- significance level

9.2 Tests about a population mean

9.4 P-values

9.3 Tests concerning a population proportion

9.5 Selecting a test procedure

Hypothesis testing

Hypothesis testing

In any hypothesis-testing problem, there are two contradictory hypotheses under consideration

- The null hypothesis, denoted by H_0 , is the claim that is initially assumed to be true
- The alternative hypothesis, denoted by H_a , is the assertion that is contradictory to H_0 .

Implicit rules (of this chapter)

- H_0 will always be stated as an equality claim.
- If θ denotes the parameter of interest, the null hypothesis will have the form

$$H_0 : \theta = \theta_0$$

- θ_0 is a specified number called the *null value*
- The alternative hypothesis will be either:
 - $H_a : \theta > \theta_0$
 - $H_a : \theta < \theta_0$
 - $H_a : \theta \neq \theta_0$

A test procedure is specified by the following:

- A test statistic T : a function of the sample data on which the decision (reject H_0 or do not reject H_0) is to be based
- A rejection region \mathcal{R} : the set of all test statistic values for which H_0 will be rejected

The null hypothesis will then be rejected if and only if the observed or computed test statistic value falls in the rejection region, i.e., $T \in \mathcal{R}$

Type I and Type II errors

- A type I error consists of rejecting the null hypothesis H_0 when it is true
- A type II error involves not rejecting H_0 when H_0 is false.

Example 2: Type I error

- It is believed that drying times of an additive follows $\mathcal{N}(\mu, 9^2)$.
- Test of hypotheses:

$$H_0 : \mu = 75$$

$$H_a : \mu < 75$$

- $n = 25$. Rule: If $\bar{x} \leq 72$, reject H_0 .
- Question: What is the probability of type I error?

$$\begin{aligned}\alpha &= P[\text{Type I error}] \\ &= P[H_0 \text{ is rejected while it is true}] \\ &= P[\bar{X} \leq 72 \text{ while } \mu = 75] \\ &= P[\bar{X} \leq 72 \text{ while } \bar{X} \sim \mathcal{N}(75, 1.8^2)] = 0.0475\end{aligned}$$

Example 2: Type II error

- It is believed that drying times of an additive follows $\mathcal{N}(\mu, 9^2)$.
- Test of hypotheses:

$$H_0 : \mu = 75$$

$$H_a : \mu < 75$$

- Experimental data is to consist of drying times from $n = 25$ test specimens: X_1, X_2, \dots, X_{25} .
- New rule:
 - Compute \bar{x}
 - If $\bar{x} \leq 72$, reject H_0 .
- Question: What are $\beta(70)$?

$\Phi(z)$

| z | .00 | .01 | .02 | .03 | .04 | .05 | .06 | .07 | .08 | .09 |
|-----|-------|-------|-------|-------|-------|-------|-------|-------|-------|-------|
| 0.0 | .5000 | .5040 | .5080 | .5120 | .5160 | .5199 | .5239 | .5279 | .5319 | .5359 |
| 0.1 | .5398 | .5438 | .5478 | .5517 | .5557 | .5596 | .5636 | .5675 | .5714 | .5753 |
| 0.2 | .5793 | .5832 | .5871 | .5910 | .5948 | .5987 | .6026 | .6064 | .6103 | .6141 |
| 0.3 | .6179 | .6217 | .6255 | .6293 | .6331 | .6368 | .6406 | .6443 | .6480 | .6517 |
| 0.4 | .6554 | .6591 | .6628 | .6664 | .6700 | .6736 | .6772 | .6808 | .6844 | .6879 |
| 0.5 | .6915 | .6950 | .6985 | .7019 | .7054 | .7088 | .7123 | .7157 | .7190 | .7224 |
| 0.6 | .7257 | .7291 | .7324 | .7357 | .7389 | .7422 | .7454 | .7486 | .7517 | .7549 |
| 0.7 | .7580 | .7611 | .7642 | .7673 | .7704 | .7734 | .7764 | .7794 | .7823 | .7852 |
| 0.8 | .7881 | .7910 | .7939 | .7967 | .7995 | .8023 | .8051 | .8078 | .8106 | .8133 |
| 0.9 | .8159 | .8186 | .8212 | .8238 | .8264 | .8289 | .8315 | .8340 | .8365 | .8389 |
| 1.0 | .8413 | .8438 | .8461 | .8485 | .8508 | .8531 | .8554 | .8577 | .8599 | .8621 |
| 1.1 | .8643 | .8665 | .8686 | .8708 | .8729 | .8749 | .8770 | .8790 | .8810 | .8830 |
| 1.2 | .8849 | .8869 | .8888 | .8907 | .8925 | .8944 | .8962 | .8980 | .8997 | .9015 |
| 1.3 | .9032 | .9049 | .9066 | .9082 | .9099 | .9115 | .9131 | .9147 | .9162 | .9177 |
| 1.4 | .9192 | .9207 | .9222 | .9236 | .9251 | .9265 | .9278 | .9292 | .9306 | .9319 |
| 1.5 | .9332 | .9345 | .9357 | .9370 | .9382 | .9394 | .9406 | .9418 | .9429 | .9441 |
| 1.6 | .9452 | .9463 | .9474 | .9484 | .9495 | .9505 | .9515 | .9525 | .9535 | .9545 |
| 1.7 | .9554 | .9564 | .9573 | .9582 | .9591 | .9599 | .9608 | .9616 | .9625 | .9633 |
| 1.8 | .9641 | .9649 | .9656 | .9664 | .9671 | .9678 | .9686 | .9693 | .9699 | .9706 |
| 1.9 | .9713 | .9719 | .9726 | .9732 | .9738 | .9744 | .9750 | .9756 | .9761 | .9767 |
| 2.0 | .9772 | .9778 | .9783 | .9788 | .9793 | .9798 | .9803 | .9808 | .9812 | .9817 |
| 2.1 | .9821 | .9826 | .9830 | .9834 | .9838 | .9842 | .9846 | .9850 | .9854 | .9857 |
| 2.2 | .9861 | .9864 | .9868 | .9871 | .9875 | .9878 | .9881 | .9884 | .9887 | .9890 |
| 2.3 | .9893 | .9896 | .9898 | .9901 | .9904 | .9906 | .9909 | .9911 | .9913 | .9916 |
| 2.4 | .9918 | .9920 | .9922 | .9925 | .9927 | .9929 | .9931 | .9932 | .9934 | .9936 |
| 2.5 | .9938 | .9940 | .9941 | .9943 | .9945 | .9946 | .9948 | .9949 | .9951 | .9952 |
| 2.6 | .9953 | .9955 | .9956 | .9957 | .9959 | .9960 | .9961 | .9962 | .9963 | .9964 |
| 2.7 | .9965 | .9966 | .9967 | .9968 | .9969 | .9970 | .9971 | .9972 | .9973 | .9974 |
| 2.8 | .9974 | .9975 | .9976 | .9977 | .9977 | .9978 | .9979 | .9979 | .9980 | .9981 |
| 2.9 | .9981 | .9982 | .9982 | .9983 | .9984 | .9984 | .9985 | .9985 | .9986 | .9986 |
| 3.0 | .9987 | .9987 | .9987 | .9988 | .9988 | .9989 | .9989 | .9989 | .9990 | .9990 |
| 3.1 | .9990 | .9991 | .9991 | .9991 | .9992 | .9992 | .9992 | .9992 | .9993 | .9993 |
| 3.2 | .9993 | .9993 | .9994 | .9994 | .9994 | .9994 | .9994 | .9995 | .9995 | .9995 |
| 3.3 | .9995 | .9995 | .9995 | .9996 | .9996 | .9996 | .9996 | .9996 | .9996 | .9997 |

- Test of hypotheses:

$$H_0 : \mu = 75$$

$$H_a : \mu < 75$$

- $n = 25$. Rule: If $\bar{x} \leq 72$, reject H_0 .

$$\begin{aligned}\beta(70) &= P[\text{Type II error when } \mu = 70] \\ &= P[H_0 \text{ is not rejected while it is false because } \mu = 70] \\ &= P[\bar{X} > 72 \text{ while } \mu = 70] \\ &= P[\bar{X} > 72 \text{ while } \bar{X} \sim \mathcal{N}(70, 1.8^2)] = 0.1335\end{aligned}$$

Example 2b

- Test of hypotheses:

$$H_0 : \mu = 75$$

$$H_a : \mu < 75$$

- $n = 25$. New rule: If $\bar{x} \leq c$, reject H_0 .
- Find the value of c to make the probability of making Type I error equal to 0.1

$$\begin{aligned}\alpha &= P[\text{Type I error}] \\ &= P[H_0 \text{ is rejected while it is true}] \\ &= P[\bar{X} \leq c \text{ while } \bar{X} \sim \mathcal{N}(75, 1.8^2)] \\ &= P\left[\frac{\bar{X} - 75}{1.8} \leq \frac{c - 75}{1.8}\right]\end{aligned}$$

- Rejection rule: $\bar{x} \leq 75 - 1.8z_\alpha$
- To make it simpler, define $z = (\bar{x} - 75)/(1.8)$, then the rule is

$$z \leq -z_\alpha$$

Example 2c

- If we want to test

$$H_0 : \mu = 75$$

$$H_a : \mu \neq 75$$

- $n = 25$. Rule: If

$$\bar{x} \leq a \quad \text{or} \quad \bar{x} \geq b$$

reject H_0 .

- Find the value of a, b to make $\alpha = 0.1$

Proposition

Suppose an experiment and a sample size are fixed and a test statistic is chosen. Then decreasing the size of the rejection region to obtain a smaller value of α results in a larger value of β for any particular parameter value consistent with H_a .

The approach adhered to by most statistical practitioners is

- specify the largest value of α that can be tolerated
- find a rejection region having that value of α rather than anything smaller
- α : the *significance level* of the test
- the corresponding test procedure is called a *level α test*

Hypothesis testing for one parameter

- 1 Identify the parameter of interest
- 2 Determine the null value and state the null hypothesis
- 3 State the appropriate alternative hypothesis
- 4 Give the formula for the test statistic
- 5 State the rejection region for the selected significance level α
- 6 Compute statistic value from data
- 7 Decide whether H_0 should be rejected and state this conclusion in the problem context

Normal population with known σ

Test about a population mean

- Null hypothesis

$$H_0 : \mu = \mu_0$$

- The alternative hypothesis will be either:
 - $H_a : \mu > \mu_0$
 - $H_a : \mu < \mu_0$
 - $H_a : \mu \neq \mu_0$

Normal population with known σ

Null hypothesis: $\mu = \mu_0$

Test statistic:

$$Z = \frac{\bar{X} - \mu_0}{\sigma/\sqrt{n}}$$

...

Alternative Hypothesis

$$H_a: \mu > \mu_0$$

$$H_a: \mu < \mu_0$$

$$H_a: \mu \neq \mu_0$$

Rejection Region for Level α Test

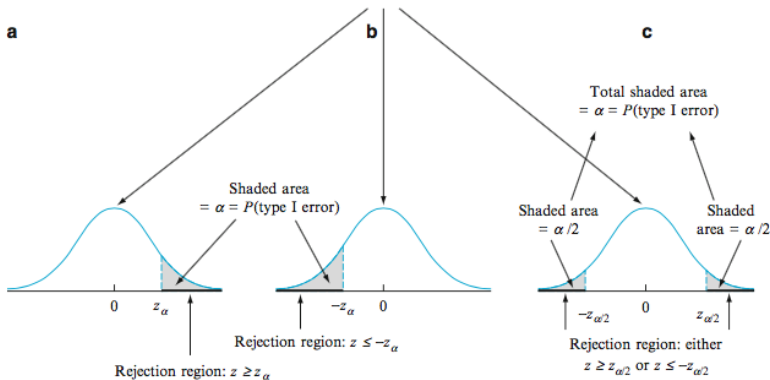
$$z \geq z_\alpha \text{ (upper-tailed test)}$$

$$z \leq -z_\alpha \text{ (lower-tailed test)}$$

$$\text{either } z \geq z_{\alpha/2} \text{ or } z \leq -z_{\alpha/2} \text{ (two-tailed test)}$$

General rule

z curve (probability distribution of test statistic Z when H_0 is true)



Example

Problem

A manufacturer of sprinkler systems used for fire protection in office buildings claims that the true average system-activation temperature is $130^{\circ}F$. A sample of $n = 9$ systems, when tested, yields a sample average activation temperature of $131.08^{\circ}F$.

If the distribution of activation times is normal with standard deviation $1.5^{\circ}F$, does the data contradict the manufacturer's claim at significance level $\alpha = 0.01$?

- Parameter of interest: $\mu =$ true average activation temperature
- Hypotheses

$$H_0 : \mu = 130$$

$$H_a : \mu \neq 130$$

- Test statistic:

$$z = \frac{\bar{x} - 130}{1.5/\sqrt{n}}$$

- Rejection region: either $z \leq -z_{0.005}$ or $z \geq z_{0.005} = 2.58$
- Substituting $\bar{x} = 131.08$, $n = 25 \rightarrow z = 2.16$.
- Note that $-2.58 < 2.16 < 2.58$. We fail to reject H_0 at significance level 0.01.
- The data does not give strong support to the claim that the true average differs from the design value.