MATH 450: Mathematical statistics

Vu Dinh

February 15th, 2019

Lecture 3: Simulations with R

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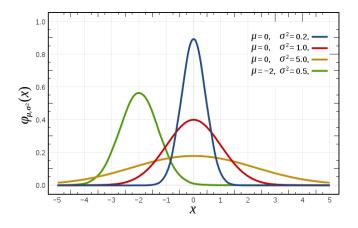
Normal random variables

Reading: 4.3

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 $\mathcal{N}(\mu, \sigma^2)$



 $E(X) = \mu$, $Var(X) = \sigma^2$

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 $\mathcal{N}(\mu, \sigma^2)$

•
$$E(X) = \mu$$
, $Var(X) = \sigma^2$

Density function

$$f(x,\mu,\sigma) = \frac{1}{\sqrt{2\pi\sigma^2}} e^{-\frac{(x-\mu)^2}{2\sigma^2}}$$

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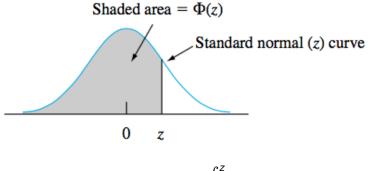
 If Z is a normal random variable with parameters μ = 0 and σ = 1, then the pdf of Z is

$$f(z)=\frac{1}{\sqrt{2\pi}}e^{-\frac{z^2}{2}}$$

and Z is called the standard normal distribution
E(Z) = 0, Var(Z) = 1

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 $\Phi(z)$



$$\Phi(z) = P(Z \le z) = \int_{-\infty}^{z} f(y) \, dy$$

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z	.00	.01	.02	.03	.04	.05	.06	.07	.08	.09
0.0	.5000	.5040	.5080	.5120	.5160	.5199	.5239	.5279	.5319	.5359
0.1	.5398	.5438	.5478	.5517	.5557	.5596	.5636	.5675	.5714	.5753
0.2	.5793	.5832	.5871	.5910	.5948	.5987	.6026	.6064	.6103	.6141
0.3	.6179	.6217	.6255	.6293	.6331	.6368	.6406	.6443	.6480	.6517
0.4	.6554	.6591	.6628	.6664	.6700	.6736	.6772	.6808	.6844	.6879
0.5	.6915	.6950	.6985	.7019	.7054	.7088	.7123	.7157	.7190	.7224
0.6	.7257	.7291	.7324	.7357	.7389	.7422	.7454	.7486	.7517	.7549
0.7	.7580	.7611	.7642	.7673	.7704	.7734	.7764	.7794	.7823	.7852
0.8	.7881	.7910	.7939	.7967	.7995	.8023	.8051	.8078	.8106	.8133
0.9	.8159	.8186	.8212	.8238	.8264	.8289	.8315	.8340	.8365	.8389
1.0	.8413	.8438	.8461	.8485	.8508	.8531	.8554	.8577	.8599	.8621
1.1	.8643	.8665	.8686	.8708	.8729	.8749	.8770	.8790	.8810	.8830
1.2	.8849	.8869	.8888	.8907	.8925	.8944	.8962	.8980	.8997	.9015
1.3	.9032	.9049	.9066	.9082	.9099	.9115	.9131	.9147	.9162	.9177
1.4	.9192	.9207	.9222	.9236	.9251	.9265	.9278	.9292	.9306	.9319
1.5	.9332	.9345	.9357	.9370	.9382	.9394	.9406	.9418	.9429	.9441
1.6	.9452	.9463	.9474	.9484	.9495	.9505	.9515	.9525	.9535	.9545
1.7	.9554	.9564	.9573	.9582	.9591	.9599	.9608	.9616	.9625	.9633
1.8	.9641	.9649	.9656	.9664	.9671	.9678	.9686	.9693	.9699	.9706
1.9	.9713	.9719	.9726	.9732	.9738	.9744	.9750	.9756	.9761	.9767

Table A.3 Standard Normal Curve Areas (cont.)

 $\Phi(z) = P(Z \le z)$

Problem

Let X be a normal random variable with mean μ and standard deviation σ .

Then

$$Z = \frac{X - \mu}{\sigma}$$

ôŔřŹ follows the standard normal distribution.

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Shifting and scaling normal random variables

If X has a normal distribution with mean μ and standard deviation σ , then

$$Z = \frac{X - \mu}{\sigma}$$

has a standard normal distribution. Thus

$$P(a \le X \le b) = P\left(\frac{a-\mu}{\sigma} \le Z \le \frac{b-\mu}{\sigma}\right)$$
$$= \Phi\left(\frac{b-\mu}{\sigma}\right) - \Phi\left(\frac{a-\mu}{\sigma}\right)$$
$$P(X \le a) = \Phi\left(\frac{a-\mu}{\sigma}\right) \quad P(X \ge b) = 1 - \Phi\left(\frac{b-\mu}{\sigma}\right)$$

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Problem

Let X be a $\mathcal{N}(3,9)$ random variable. Compute $P[X \le 5.25]$.

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Linear combination of normal random variables

Theorem

Let $X_1, X_2, ..., X_n$ be independent normal random variables (with possibly different means and/or variances). Then

$$T = a_1 X_1 + a_2 X_2 + \ldots + a_n X_n$$

ôŔřŹ also follows the normal distribution.

Theorem

Let $X_1, X_2, ..., X_n$ be independent normal random variables (with possibly different means and/or variances). Then

$$T = a_1 X_1 + a_2 X_2 + \ldots + a_n X_n$$

ôŔřŹ also follows the normal distribution.

What are the mean and the standard deviation of *T*?

•
$$E(T) = a_1 E(X_1) + a_2 E(X_2) + \dots + a_n E(X_n)$$

• $\sigma_T^2 = a_1^2 \sigma_{X_1}^2 + a_2^2 \sigma_{X_2}^2 + \dots + a_n^2 \sigma_{X_n}^2$

Problem

Assume that

$$X_1 \sim \mathcal{N}(10,9)$$
 and $X_2 \sim \mathcal{N}(30,16)$

are independent.

What is the distribution of $X_1 - X_2$?

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Problem

A concert has three pieces of music to be played before intermission. The time taken to play each piece has a normal distribution.

Assume that the three times are independent of each other. The mean times are 15, 30, and 20 min, respectively, and the standard deviations are 1, 2, and 1.5 min, respectively.

What is the distribution of the length of the concert?

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Working with R

Reading: 4.3

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• manually create a vector a with entry values

$$a = c(1,2,6,8,5,3,-1,2,1,0)$$

create a zero vector with length n = 25

$$a = rep(0, 25)$$

- *a*[*i*] is the *i*th element of *a*
- manipulate all entries at the same time using 'for' loop

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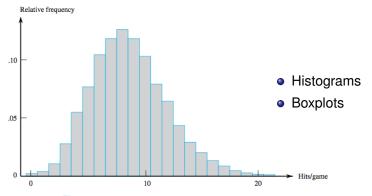


Figure 1.6 Histogram of number of hits per nine-inning game

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- The Mean
- The Median
- Trimmed Means

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The sample mean \overline{x} of observations x_1, x_2, \ldots, x_n is given by

$$\bar{x} = \frac{x_1 + x_2 + \dots + x_n}{n} = \frac{\sum_{i=1}^n x_i}{n}$$

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Step 1: ordering the observations from smallest to largest

$$\widetilde{x} = \begin{cases} \text{The single} \\ \text{middle} \\ \text{value if } n \\ \text{is odd} \end{cases} = \left(\frac{n+1}{2}\right)^{\text{th}} \text{ordered value} \\ \text{The average} \\ \text{of the two} \\ \text{middle} \\ \text{values if } n \\ \text{is even} \end{cases} = \text{average of } \left(\frac{n}{2}\right)^{\text{th}} \text{ and } \left(\frac{n}{2}+1\right)^{\text{th}} \text{ordered values} \end{cases}$$

Median is not affected by outliers

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Measures of locations: trimmed mean

- A α % trimmed mean is computed by:
 - eliminating the smallest α % and the largest α % of the sample
 - averaging what remains
- $\alpha = 0 \rightarrow$ the mean
- $\alpha \approx 50 \rightarrow$ the median

The sample variance, denoted by s^2 , is given by

$$s^{2} = \frac{\sum (x_{i} - \bar{x})^{2}}{n - 1} = \frac{S_{xx}}{n - 1}$$

The **sample standard deviation**, denoted by *s*, is the (positive) square root of the variance:

$$s = \sqrt{s^2}$$

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The **sample standard deviation**, denoted by *s*, is the (positive) square root of the variance:

$$s = \sqrt{s^2}$$

- Why squared? Because it is easier to do math with x^2 than |x|
- Why (n−1)? Because that makes s² an <u>unbiased estimator</u> of the population variance (Chapter 7)

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Computing formula for s^2

$$S_{xx} = \sum (x_i - \bar{x})^2 = \sum x_i^2 - \frac{\left(\sum x_i\right)^2}{n}$$

Proof Because $\bar{x} = \sum x_i/n$, $n\bar{x}^2 = (\sum x_i)^2/n$. Then,

$$\sum (x_i - \bar{x})^2 = \sum (x_i^2 - 2\bar{x} \cdot x_i + \bar{x}^2) = \sum x_i^2 - 2\bar{x} \sum x_i + \sum (\bar{x})^2$$
$$= \sum x_i^2 - 2\bar{x} \cdot n\bar{x} + n(\bar{x})^2 = \sum x_i^2 - n(\bar{x})^2$$

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Let x_1, x_2, \ldots, x_n be a sample and c be a constant.

1. If
$$y_1 = x_1 + c$$
, $y_2 = x_2 + c$, ..., $y_n = x_n + c$, then $s_y^2 = s_x^2$, and
2. If $y_1 = cx_1, \ldots, y_n = cx_n$, then $s_y^2 = c^2 s_x^2$, $s_y = |c| s_x$,

where s_x^2 is the sample variance of the x's and s_y^2 is the sample variance of the y's.

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Order the *n* observations from smallest to largest and separate the smallest half from the largest half; the median \tilde{x} is included in both halves if *n* is odd. Then the **lower fourth** is the median of the smallest half and the **upper fourth** is the median of the largest half. A measure of spread that is resistant to outliers is the **fourth spread** f_s , given by

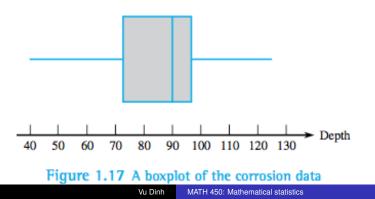
 $f_s =$ upper fourth – lower fourth

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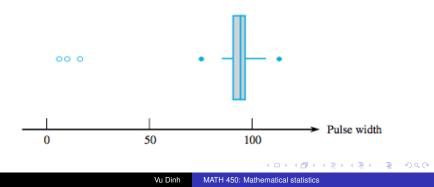
40 52 55 60 70 75 85 85 90 90 92 94 94 95 98 100 115 125 125

The five-number summary is as follows:

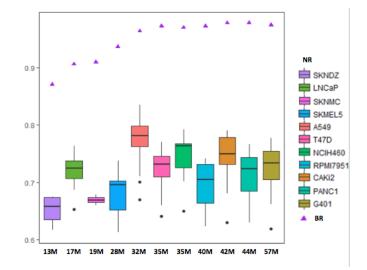
smallest $x_i = 40$ lower fourth = 72.5 $\tilde{x} = 90$ upper fourth = 96.5 largest $x_i = 125$



Any observation farther than $1.5f_s$ from the closest fourth is an **outlier**. An outlier is **extreme** if it is more than $3f_s$ from the nearest fourth, and it is **mild** otherwise.



Comparative boxplots



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- Experiment: throw a fair die 2 times
- Before the experiment, denote the random variables that describes the outcome of the first throw and the second throws by X₁ and X₂, respectively
- X₁ and X₂ have the same probability mass function

$$p(x) = 1/6, x = 1, 2, 3, 4, 5, 6$$

- Do the experiment, obtain outcomes x₁ and x₂
- x₁ may be different from x₂

Definition

The random variables $X_1, X_2, ..., X_n$ are said to form a random sample of size *n* if

- the X_i 's are independent
- every X_i has the same probability distribution

Each of the X_i 's is called an instance, or a realization of the distribution.

Sometimes, people refer to X_i as copies of the same random variable X

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Given a probability distribution, one want to create a sample of it

- in real life: statistical sampling
- e using computer: simulation

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• the uniform distribution on (a, b) has density

$$f(x) = \begin{cases} \frac{1}{b-a} & \text{if } x \in (a,b) \\ 0 & \text{elsewhere} \end{cases}$$

• To generate the uniform distribution on (0,1), use the function *runif*

$$b = runif(200)$$

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Assume that we want to simulate a Bernoulli random variable

$$p(x) = \begin{cases} 0.6 & \text{if } x = 0\\ 0.4 & \text{if } x = 1\\ 0 & \text{otherwise} \end{cases}$$

- Step 1: generate *u* from the uniform distribution on (0,1)
- Step 2: If u < 0.4, then set x = 1; otherwise set x = 0

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 Question: How to simulate samples from the following distribution

$$p(x) = \begin{cases} 0.2 & \text{if } x = 3\\ 0.3 & \text{if } x = 5\\ 0.5 & \text{if } x = 7\\ 0 & \text{otherwise} \end{cases}$$

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Simulate continuous random variables

 Question: How to simulate samples from the following distribution

$$f(x) = \begin{cases} 2e^{-2x} & \text{if } x > 0\\ 0 & \text{otherwise} \end{cases}$$

• The distribution function of X is

$$F(x) = \begin{cases} 1 - e^{-2x} & \text{if } x > 0\\ 0 & \text{otherwise} \end{cases}$$

- Step 1: generate *u* from the uniform distribution on (0,1)
- Step 2: Solve equation F(x) = u
- Set x as the solution

$$x=-\frac{1}{2}\ln(1-u)$$

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Theorem

Let X be a continuous random variable with probability distribution function F. Then F(X) is a uniform random variable over (0,1).

Proof.

Let
$$Y = F(X)$$
, then $Y \in [0, 1]$ and for all $y \in (0, 1)$

$$F_{Y}(y) = P[Y \le y] = P[F(X) \le y] = P[X \le F^{-1}(y)] = F(F^{-1}(y)) = y$$

thus

$$f_{Y}(y) = \begin{cases} 1 & \text{if } y \in (0,1) \\ 0 & \text{otherwise} \end{cases}$$

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