# Mathematical statistics 

February $25^{\text {th }}, 2018$
Lecture 6: Statistics and sampling distribution (cont.)

| Week 1 | Probability reviews |
| :---: | :---: |
|  | Chapter 6: Statistics and Sampling Distributions |
| Week 4 | Chapter 7: Point Estimation |
| Week 7 | Chapter 8: Confidence Intervals |
| Week 10 | Chapter 9: Test of Hypothesis |
| Week 14 | Regression |

## Overview

6.1 Statistics and their distributions
6.2 The distribution of the sample mean
6.3 The distribution of a linear combination

Order $6.1 \rightarrow 6.3 \rightarrow 6.2$

## Random sample



## Definition

The random variables $X_{1}, X_{2}, \ldots, X_{n}$ are said to form a (simple) random sample of size $n$ if
(1) the $X_{i}$ 's are independent random variables
(2) every $X_{i}$ has the same probability distribution

## Questions for this chapter

Given a random sample $X_{1}, X_{2}, \ldots, X_{n}$, and

$$
T=a_{1} X_{1}+a_{2} X_{2}+\ldots+a_{n} X_{n}
$$

- If we know the distribution of $X_{i}$ 's, can we obtain the distribution of $T$ ?
- If we don't know the distribution of $X_{i}$ 's, can we still obtain/approximate the distribution of $T$ ?


## 6.1: Summary

(1) If the distribution and the statistic $T$ is simple, try to construct the pmf of the statistic (as in Example 1)
(2) If the probability density function $f_{X}(x)$ of $X$ 's is known, the

- try to represent/compute the cumulative distribution (cdf) of $T$

$$
\mathbb{P}[T \leq t]
$$

- take the derivative of the function (with respect to $t$ )


## 6.2: Linear combination of normal random variables

## Theorem

Let $X_{1}, X_{2}, \ldots, X_{n}$ be independent normal random variables (with possibly different means and/or variances). Then

$$
T=a_{1} X_{1}+a_{2} X_{2}+\ldots+a_{n} X_{n}
$$

also follows the normal distribution.
What are the mean and the standard deviation of $T$ ?

- $E(T)=a_{1} E\left(X_{1}\right)+a_{2} E\left(X_{2}\right)+\ldots+a_{n} E\left(X_{n}\right)$
- $\sigma_{T}^{2}=a_{1}^{2} \sigma_{X_{1}}^{2}+a_{2}^{2} \sigma_{X_{2}}^{2}+\ldots+a_{n}^{2} \sigma_{X_{n}}^{2}$


## Moment generating function

## Definition

The moment generating function (mgf) of a continuous random variable $X$ is

$$
M_{X}(t)=E\left(e^{t X}\right)=\int_{-\infty}^{\infty} e^{t x} f_{X}(x) d x
$$

Reading: 3.4 and 4.2

## Moment generating function

## Property

Two distributions have the same pdf if and only if they have the same moment generating function

## Moment generating function

| Distribution | Moment-generating function $M_{X}(t)$ |
| :---: | :---: |
| Bernoulli $P(X=1)=p$ | $1-p+p e^{t}$ |
| Geometric $(1-p)^{k-1} p$ | $\begin{gathered} \frac{p e^{t}}{1-(1-p) e^{t}} \\ \forall t<-\ln (1-p) \end{gathered}$ |
| Binomial $\mathrm{B}(n, p)$ | $\left(1-p+p e^{t}\right)^{n}$ |
| Poisson Pois ( $\lambda$ ) | $e^{\lambda\left(e^{t}-1\right)}$ |
| Uniform (continuous) $\mathrm{U}(\mathrm{a}, \mathrm{b})$ | $\frac{e^{t b}-e^{t a}}{t(b-a)}$ |
| Uniform (discrete) $\mathrm{U}(\mathrm{a}, \mathrm{b})$ | $\frac{e^{a t}-e^{(b+1) t}}{(b-a+1)\left(1-e^{t}\right)}$ |
| Normal $\boldsymbol{N}\left(\mu, \sigma^{2}\right)$ | $e^{t \mu+\frac{1}{2} \sigma^{2} t^{2}}$ |
| Chi-squared $\chi_{k}^{2}$ | $(1-2 t)^{-\frac{k}{2}}$ |
| Gamma $\Gamma(k, \theta)$ | $(1-t \theta)^{-k} ; \forall t<\frac{1}{\theta}$ |
| Exponential $\operatorname{Exp}(\lambda)$ | $\left(1-t \lambda^{-1}\right)^{-1},(t<\lambda)$ |

[^0]
## Moment generating function

## Definition

Let $X_{1}, X_{2}$ be a 2 independent random variables and $T=X_{1}+X_{2}$, then

$$
M_{T}(t)=M_{X_{1}}(t) M_{X_{2}}(t)
$$

Hint:

$$
M_{T}(t)=E\left(e^{t T}\right)=E\left(e^{t\left(X_{1}+X_{2}\right)}\right)=E\left(e^{t X_{1}} \cdot e^{t X_{2}}\right)
$$

## Example 3

## Problem

Given that the mgf of a Poisson variables with mean $\lambda$ is

$$
e^{\lambda\left(e^{t}-1\right)}
$$

Suppose $X$ and $Y$ are independent Poisson random variables, where $X$ has mean a and $Y$ has mean $b$. Show that $T=X+Y$ also follows the Poisson distribution.

## Example 4

## Problem

Given that the mgf of a normal random variables with mean $\mu$ and variance $\sigma^{2}$ is

$$
e^{\mu t+\frac{\sigma^{2}}{2} t^{2}}
$$

Suppose $X$ and $Y$ are independent normal random variables. Show that $T=X+Y$ also follows the normal distribution.

## Shaded area $=\Phi(z)$



Table A. 3 Standard Normal Curve Areas (cont.) $\quad \Phi(z)=P(Z \leq z)$

| $\boldsymbol{z}$ | $\mathbf{. 0 0}$ | $\mathbf{. 0 1}$ | $\mathbf{. 0 2}$ | $\mathbf{. 0 3}$ | $\mathbf{. 0 4}$ | $\mathbf{. 0 5}$ | $\mathbf{. 0 6}$ | $\mathbf{. 0 7}$ | $\mathbf{. 0 8}$ | $\mathbf{. 0 9}$ |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 0.0 | .5000 | .5040 | .5080 | .5120 | .5160 | .5199 | .5239 | .5279 | .5319 | .5359 |
| 0.1 | .5398 | .5438 | .5478 | .5517 | .5557 | .5596 | .5636 | .5675 | .5714 | .5753 |
| 0.2 | .5793 | .5832 | .5871 | .5910 | .5948 | .5987 | .6026 | .6064 | .6103 | .6141 |
| 0.3 | .6179 | .6217 | .6255 | .6293 | .6331 | .6368 | .6406 | .6443 | .6480 | .6517 |
| 0.4 | .6554 | .6591 | .6628 | .6664 | .6700 | .6736 | .6772 | .6808 | .6844 | .6879 |
| 0.5 | .6915 | .6950 | .6985 | .7019 | .7054 | .7088 | .7123 | .7157 | .7190 | .7224 |
| 0.6 | .7257 | .7291 | .7324 | .7357 | .7389 | .7422 | .7454 | .7486 | .7517 | .7549 |
| 0.7 | .7580 | .7611 | .7642 | .7673 | .7704 | .7734 | .7764 | .7794 | .7823 | .7852 |
| 0.8 | .7881 | .7910 | .7939 | .7967 | .7995 | .8023 | .8051 | .8078 | .8106 | .8133 |
| 0.9 | .8159 | .8186 | .8212 | .8238 | .8264 | .8289 | .8315 | .8340 | .8365 | .8389 |
| 1.0 | .8413 | .8438 | .8461 | .8485 | .8508 | .8531 | .8554 | .8577 | .8599 | .8621 |
| 1.1 | .8643 | .8665 | .8686 | .8708 | .8729 | .8749 | .8770 | .8790 | .8810 | .8830 |
| 1.2 | .8849 | .8869 | .8888 | .8907 | .8925 | .8944 | .8962 | .8980 | .8997 | .9015 |
| 1.3 | .9032 | .9049 | .9066 | .9082 | .9099 | .9115 | .9131 | .9147 | .9162 | .9177 |
| 1.4 | .9192 | .9207 | .9222 | .9236 | .9251 | .9265 | .9278 | .9292 | .9306 | .9319 |
| 1.5 | .9332 | .9345 | .9357 | .9370 | .9382 | .9394 | .9406 | .9418 | .9429 | .9441 |
| 1.6 | .9452 | .9463 | .9474 | .9484 | .9495 | .9505 | .9515 | .9525 | .9535 | .9545 |
| 1.7 | .9554 | .9564 | .9573 | .9582 | .9591 | .9599 | .9608 | .9616 | .9625 | .9633 |
| 1.8 | .9641 | .9649 | .9656 | .9664 | .9671 | .9678 | .9686 | .9693 | .9699 | .9706 |
| 1.9 | .9713 | .9719 | .9726 | .9732 | .9738 | .9744 | .9750 | .9756 | .9761 | .9767 |

## Example 1

## Problem

Let $X_{1}, X_{2}, \ldots, X_{16}$ be a random sample from $\mathcal{N}(1,4)$ (that is, normal distribution with mean $\mu=1$ and standard deviation $\sigma=2$ ).
Let $\bar{X}$ be the sample mean

$$
\bar{X}=\frac{X_{1}+X_{2}+\ldots+X_{16}}{16}
$$

- What is the distribution of $\bar{X}$ ?
- Compute $P[\bar{X} \leq 1.82]$


## Example 1*

## Problem

Let $X_{1}, X_{2}, \ldots, X_{n}$ be a random sample from $\mathcal{N}\left(\mu, \sigma^{2}\right)$ (that is, normal distribution with mean $\mu$ and standard deviation $\sigma$ ). Let $\bar{X}$ be the sample mean

$$
\bar{X}=\frac{X_{1}+X_{2}+\ldots+X_{n}}{n}
$$

What is the distribution of $\bar{X}$ ?


[^0]:    Mathematical statistics

