Mathematical techniques in data science

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Lecture 6: Classification - Linear Discriminant Analysis

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11	Clustering: K-means \rightarrow Spectral Clustering
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- Logistic regression
- Linear Discriminant Analysis
- Support Vector Machines
- Nearest neighbours

Classification: Logistic regression

- Suppose we work with binary outputs $\mathcal{Y}=\{0,1\},$ and \mathcal{X} is a subset of \mathbb{R}^d
- Goal: Given input X, we want to model the probability that Y = 1

$$P[Y=1|X=x]$$

Logistic function and logit function

Transformation between $(-\infty,\infty)$ and [0,1]





 $logit(p) = \log \frac{p}{1-p}$

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Assumption

Given X = x, Y is a Bernoulli random variable with parameter p(x) = P[Y = 1|X = x] and

$$logit(p(x)) = \log \frac{p(x)}{1 - p(x)} = \log \frac{P[Y = 1 | X = x]}{P[Y = 0 | X = x]} = x^{T} \beta$$

for some vector $\beta \in \mathbb{R}^{d+1}$.

Note: Here we denote

$$x^{\mathsf{T}}\beta = \beta_0 + \beta_1 x_1 + \ldots + \beta_d x_d$$

Logistic regression: Assumptions



• Remember that for Bernoulli r.v. with parameter p

$$P[Y = y] = p^{y}(1-p)^{1-y}, y \in \{0,1\}$$

• Given samples $(x_1, y_1), (x_2, y_2), ..., (x_n, y_n)$, we have

$$L(\beta) = \prod_{i=1}^{n} p(x_i, \beta)^{y_i} (1 - p(x_i, \beta))^{1 - y_i}$$

• Maximum likelihood (ML): maximize this likelihood function

- The optimization needs to be performed by a numerical optimization method
- Penalties can be added to regularize the problem to avoid overfitting

$$\min_{\beta} -\ell(\beta) + \alpha \|\beta\|_1$$

or

$$\min_{\beta} -\ell(\beta) + \alpha \|\beta\|_2$$

Logistic regression with more than 2 classes

- \bullet Suppose now the response can take any of $\{1,\ldots,K\}$ values
- We use the categorical distribution instead of the Bernoulli distribution

$$P[Y = k | X = x] = p_k(x), \quad \sum_{k=1}^{K} p_k(x) = 1.$$

Model

$$p_k(x) = rac{e^{x^T \beta^{(k)}}}{\sum_{k=1}^{K} e^{x^T \beta^{(k)}}}$$

Logistic regression: Assumptions



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Feed-forward neural network



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Classification: Linear discriminant analysis

- Suppose we work with outputs $\mathcal{Y} = \{1, 2, \dots, K\}$, and \mathcal{X} is a subset of \mathbb{R}^d
- Goal: Given input X, we want to model the probability that Y condition on X

$$P[Y=i|X=x], \quad i\in\mathcal{Y}$$

• But some time, P[X = x | Y = i] is easier to model!



$$P(X = x | Y = \text{red}).$$

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Bayes' formula

$$P(A|B) = \frac{P(B|A)P(A)}{P(B)}$$



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Linear discriminant analysis

Suppose

Then

$$P(Y = i | X = x) = \frac{P(X = x | Y = i)P(Y = i)}{\sum_{j=1}^{K} P(X = x | Y = j)P(Y = j)}$$
$$= \frac{f_i(x)\pi_i}{\sum_{j=1}^{K} f_j(x)\pi_j}$$

Model P(X = x | Y = i) using Gaussian distributions

The natural model for $f_i(x)$ is the multivariate Gaussian distribution

$$f_i(x) = \frac{1}{\sqrt{(2\pi)^p \det(\Sigma_i)}} e^{-\frac{1}{2}(x-\mu_i)^T \Sigma_i^{-1}(x-\mu_i)}, \quad x \in \mathbb{R}^p$$

- μ: mean vector
- Σ: covariance matrix

$$\Sigma = E[(X - \mu)^T (X - \mu)]$$

Model P(X = x | Y = i) using Gaussian distributions



FIGURE 4.5. Two multivariate Gaussian density functions are shown, with p = 2. Left: The two predictors are uncorrelated. Right: The two variables have a correlation of 0.7.

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The natural model for $f_i(x)$ is the multivariate Gaussian distribution

$$f_i(x) = \frac{1}{\sqrt{(2\pi)^p \det(\Sigma_i)}} e^{-\frac{1}{2}(x-\mu_i)^T \Sigma_i^{-1}(x-\mu_i)}, \quad x \in \mathbb{R}^p$$

• Linear discriminant analysis (LDA): We assume

$$\Sigma_1=\Sigma_2=\ldots=\Sigma_K$$

• Quadratic discriminant analysis (QDA): general cases

We need to estimate

- An estimate of the class probabilities π_i
- Estimate the mean vectors μ_1, \ldots, μ_K
- Estimate the covariance matrices $\Sigma_1, \ldots, \Sigma_K$ (or Σ for LDA)

Parameter estimation: LDA

Suppose we have dataset $(x_1, y_1), (x_2, y_2), \dots, (x_n, y_n)$ where n_i observations have label *i*.

• An estimate of the class probabilities π_i

$$\hat{\pi}_i = \frac{n_i}{n}$$

• Estimate the mean vectors

$$\hat{\mu}_i = \frac{1}{n_i} \sum_{y_j = i} x_j$$

• Estimate the covariance matrix $\boldsymbol{\Sigma}$

$$\hat{\Sigma} = \frac{1}{N-K} \sum_{i=1}^{K} \sum_{y_j=i} (x_j - \hat{\mu}_i) (x_j - \hat{\mu}_i)^T$$

Suppose we have dataset $(x_1, y_1), (x_2, y_2), \dots, (x_n, y_n)$ where n_i observations have label *i*.

A new instance x arrives, how to predict label of x?

• Compute $\hat{\pi}_i$, $\hat{\mu}_i$ and $\hat{\Sigma}$

Compute

$$P(Y = i | X = x) \approx p_k(x) = \frac{f_i(x, \hat{\mu}_i, \hat{\Sigma}) \hat{\pi}_i}{\sum_{j=1}^K f_j(x, \hat{\mu}_j, \hat{\Sigma}) \hat{\pi}_j}$$

• Bayes classifier: assign an observation to the class for which the posterior probability $p_k(x)$ is greatest.



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LDA: linearity of the decision boundary

Note that

$$p_i(x) = \frac{f_i(x, \hat{\mu}_i, \hat{\Sigma})\hat{\pi}_i}{\sum_{j=1}^K f_j(x, \hat{\mu}_j, \hat{\Sigma})\hat{\pi}_j}$$

and

$$f_i(x) = \frac{1}{\sqrt{(2\pi)^p \det(\hat{\Sigma})}} e^{-\frac{1}{2}(x-\hat{\mu}_i)^T \Sigma^{-1}(x-\hat{\mu}_i)}, \quad x \in \mathbb{R}^p$$

Thus

$$\log \frac{p_{i}(x)}{p_{k}(x)} = \log \frac{\hat{\pi}_{i}}{\hat{\pi}_{k}} - \frac{1}{2} (\hat{\mu}_{i} + \hat{\mu}_{k})^{T} \hat{\Sigma}^{-1} (\hat{\mu}_{i} - \hat{\mu}_{k}) + x^{T} \hat{\Sigma}^{-1} (\hat{\mu}_{i} - \hat{\mu}_{k})$$
$$= \beta_{0} + x^{T} \beta$$

Recall that for logistic regression

$$\log \frac{P[Y=1|X=x]}{P[Y=0|X=x]} = \beta_0 + x^T \beta$$

- Both methods use linear decision boundary
- Both are simple, and often perform very well. However
- The probability models are different
- The estimations are different

- Estimating covariance matrices when $n \ll p$ is challenging
- The sample covariance $\hat{\Sigma}$ is singular when $n \ll p$
- Need regularization

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